# Optimally Combining Outcomes to Improve Prediction

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# Acknowledgments

#### **Collaborators:**

Mark van der Laan, Alan Hubbard, Ben Arnold, Jack Colford, Andrew Mertens, Oleg Sofyrgin, Jonathan French, Aryeh Stein, Shasha Jumbe

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## **Outline**

#### Motivation

#### Defining target parameters

- Measuring accuracy of predictions
- Optimal predictor and weights

#### Estimation

- Super learning
- ► Estimating weights

#### **Evaluating predictions**

- ► Estimation and inference
- ► Variable importance

#### Simulation

Data Analysis

Conclusions and future directions

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The observed data are n i.i.d. copies of O = (X, Y).

- ► X = D covariates
- ► Y = J standardized outcomes

Together Y represents an unmeasured outcome of interest.

- ► Inflammation in CVD
- ► Immune response in HIV
- ► Subject area tests in cognitive development

Researchers are interested in prediction of unmeasured outcome using X.

Can we use neonatal information to predict neurocognitive outcomes later in life?

► Early identification of at-risk children.

What covariates are important for prediction?

► Informs what information to collect to screen children.

## **Motivation: PROBIT**

The Promotion of Breastfeeding Intervention Trial enrolled pregnant mothers in 1996–97 (Kramer et al, 2001).

Variables
breastfeeding encouraged
household animals
age, height, weight, education, siblings
employment status
gestational age, Apgar score
WAZ, HAZ, HCAZ (0,1,2,3,6,9,12 months)
mother smoked during pregnancy, mother
drank during pregnancy
Matrix, Block, Vocabulary, Similarities
-

#### **Motivation: CLHNS**

The Cebu Longitudinal Health and Nutrition Survey enrolled pregnant mothers in 1983–84 (Feranil et al, 2008).

Group name	Variables
Health care	health care access, preventive health care
Household	child:adult ratio, child dependency ratio,
	crowding index, urban score
Socioeconomic	total income, SES
Water/sanitation	sanitation, access to clean water
Parental	mother age, father age, mother height, mother
	education (years), father education (years),
	marital status, mother age first child, parity
Growth	WAZ, HAZ (0,6,12,18,24 months)
Other	mother smoked during pregnancy, child's sex,
	gestational age at birth
Achievement tests	Math, Cebuano, English
(age 11)	<del>-</del>

How to combine test scores to measure "neurocognition"?

Give equal weight to all scores?

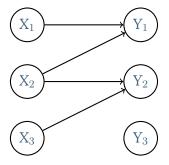
What if some scores are noisy or not related to covariates?

PCA or factor analysis to combine scores?

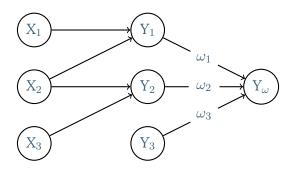
► Not related to scientific goal.

Use the combination that is predicted most accurately.

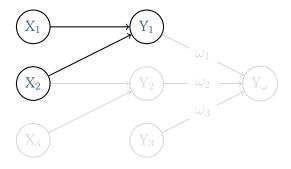
Consider this simple associative directed acyclic graph.



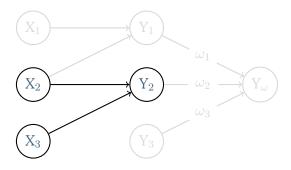
Let  $Y_{\omega} = \sum\limits_{j=1}^{J} \omega_j Y_j$ , with  $\omega_j \geq 0$  for all j and  $\sum\limits_{j=1}^{J} \omega_j = 1$ .



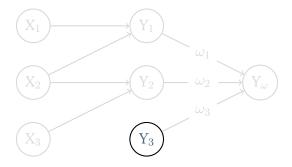
Predicting only  $Y_1$  ignores  $X_3$ 's association with Y.



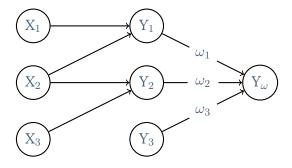
Predicting only  $Y_2$  ignores  $X_1$ 's association with Y.



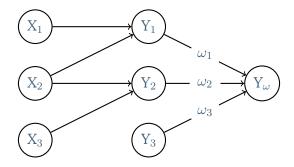
Predicting  $Y_3$  adds noise, wastes type-1 error.



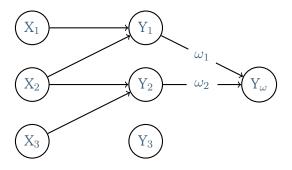
Predicting  $Y_{\omega}$  uses all of X.



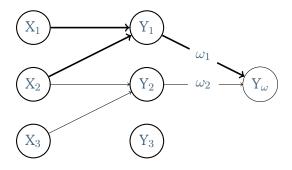
We could be clever in choosing weights if we knew the DAG.



Outcome Y<sub>3</sub> gets no weight.



Outcomes  $Y_1$  and  $Y_2$  get weight based on accuracy of predictions.



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Suppose we are given  $\omega$  and  $\psi_{\omega}: \mathcal{X} \to \mathbb{R}$  to predict  $Y_{\omega}$ .

A measure of accuracy of  $\psi_{\omega}$  is MSE:

$$\mathcal{E}_0(\psi_\omega) = E_0 \left[ \{ Y_\omega - \psi_\omega(X) \}^2 \right] \; . \label{eq:epsilon}$$

However, MSE depends on scale and variability of  $Y_{\omega}$ .

► Hard to compare across studies, outcomes.

To obtain a scale-free measure, we use a nonparametric version of  $\mathbb{R}^2$ .

$$R_{0,\omega}^{2}(\psi_{\omega}) = \frac{E_{0} \left[ \{ Y_{\omega} - \mu_{0,\omega} \}^{2} \right] - E_{0} \left[ \{ Y_{\omega} - \psi_{\omega}(X) \}^{2} \right]}{E_{0} \left[ \{ Y_{\omega} - \mu_{0,\omega} \}^{2} \right]}.$$

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$$R_{0,\omega}^2(\psi_\omega) = \frac{\overbrace{E_0\left[\left\{Y_\omega - \mu_{0,\omega}\right\}^2\right]}^{\text{MSE of } \mu_{0,\omega}} - E_0\left[\left\{Y_\omega - \psi_\omega(X)\right\}^2\right]}_{E_0\left[\left\{Y_\omega - \mu_{0,\omega}\right\}^2\right]}.$$

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# Predicting combined outcome

The criteria  $R_{0,\omega}^2$  provides a way to compare prediction functions for  $Y_{\omega}$ .

- ► Large  $R_{0,\omega}^2$  indicates accurate predictions.
- ►  $R_{0,\omega}^2 = 1$  indicates perfect predictions.
- ►  $R_{0,\omega}^2 < 0$  means predictions worse than  $\mu_{0,\omega}$ !

The maximizer over all  $\mathcal{X} \to \mathbb{R}$  is

$$\psi_{0,\omega}(X) = E_0(Y_\omega \mid X) .$$

This fact plays a key role in how we will construct a prediction function.

# Optimal weights

For any  $\omega$ , the function  $\psi_{0,\omega}$  gives the most accurate predictions of  $Y_{\omega}$ .

We also want weights that lead to most accurate predictions of combined outcome.

Formally, we define

$$\omega_0 = \operatorname{argmax}_{\omega} R_{0,\omega}^2(\psi_{0,\omega}) .$$

The statistical goal is to estimate  $\psi_{0,\omega_0}$  and  $\omega_0$ .

#### **Caveats**

The sense in which this combination is optimal is strictly related to prediction.

Not how well combined outcome measures "latent" outcome.

The optimal weights could give zero weight to some outcomes.

► These outcomes may still be important!

The procedure is best viewed as an exploratory analysis.

However, it can be fully pre-specified!

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# **Estimation**

The statistical goal is to estimate  $\psi_{0,\omega_0}$  and  $\omega_0$ .

At first glance, it looks like a difficult optimization problem,

$$\psi_{0,\omega_0} = \operatorname{argmax}_{\psi} R_{0,\omega_0}^2(\psi)$$
  
$$\omega_0 = \operatorname{argmax}_{\omega} R_{0,\omega}^2(\psi_{0,\omega})$$

The problem is made easier by recognizing

$$\psi_{0,\omega}(X) = E_0(Y_\omega \mid X) = E_0\left(\sum_{j=1}^J \omega_j Y_j \mid X\right) = \sum_{j=1}^J \omega_j E_0(Y_j \mid X) .$$

## **Estimation**

For any  $\omega$ ,  $\psi_{0,\omega}$  is weighted sum of conditional means.

This allows the optimization to be split up:

- 1. Estimate  $E_0(Y_j \mid X)$  for j = 1, ..., J.
- 2. Combine estimates  $\sum_{j=1}^{J} \omega_j \hat{E}(Y_j \mid X)$ .
- 3. Optimize over weights using estimated prediction function.

However, we must take care to avoid overfitting!

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How should we estimate  $E_0(Y_j \mid X)$  for a given j?

How should we estimate  $E_0(Y_j \mid X)$  for a given j?

► Linear regression

How should we estimate  $E_0(Y_j \mid X)$  for a given j?

► Linear regression, with interactions

How should we estimate  $E_0(Y_j \mid X)$  for a given j?

► Linear regression, with interactions, and nonlinear terms

How should we estimate  $E_0(Y_j \mid X)$  for a given j?

► Linear regression, with interactions, and nonlinear terms, or splines

How should we estimate  $E_0(Y_j \mid X)$  for a given j?

► Linear regression, with interactions, and nonlinear terms, or splines (with different degrees?)

- ► Linear regression, with interactions, and nonlinear terms, or splines (with different degrees?)
- ► Penalized linear regression

- ► Linear regression, with interactions, and nonlinear terms, or splines (with different degrees?)
- ► Penalized linear regression, with different penalties?

- ► Linear regression, with interactions, and nonlinear terms, or splines (with different degrees?)
- ► Penalized linear regression, with different penalties?
- ► Random forests

- ► Linear regression, with interactions, and nonlinear terms, or splines (with different degrees?)
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- ► Random forests, with different tuning parameters?

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- ► Penalized linear regression, with different penalties?
- ► Random forests, with different tuning parameters?
- Gradient boosting?

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- ► Penalized linear regression, with different penalties?
- Random forests, with different tuning parameters?
- Gradient boosting? Support vector machines?

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- Highly adaptive lasso?

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- Highly adaptive lasso?
- ► Variable selection?

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- ► Highly adaptive lasso?
- ► Variable selection?
- ► Ad infinitum...

How should we estimate  $E_0(Y_j \mid X)$  for a given j?

- ► Linear regression, with interactions, and nonlinear terms, or splines (with different degrees?)
- ► Penalized linear regression, with different penalties?
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- ► Highly adaptive lasso?
- ► Variable selection?
- ► Ad infinitum...

The best algorithm for estimating depends on the (unknown) truth!

► Might be different for different outcomes.

We use  $\psi : \mathcal{S} \to \Psi$  to denote an algorithm.

- ▶ S is all subsets of  $\{1, ..., n\}$ .
- $\Psi$  is set of  $\mathcal{X} \to \mathbb{R}$ .

Given a data set, an algorithm:

- 1. takes as input a subset of observations;
- 2. uses observations to create a prediction function;
- 3. returns prediction function.

We refer to this process as training an algorithm.

Say we have M algorithms that could be used to estimate  $E_0(Y_j \mid X)$ .

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How can we evaluate these algorithms?

► Train algorithms on full data, see which has largest empirical R<sup>2</sup>.

Say we have M algorithms that could be used to estimate  $E_0(Y_j \mid X)$ .

- ► Train algorithms on full data, see which has largest empirical R<sup>2</sup>.
  - Overfit!

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- ► Train algorithms on full data, see which has largest empirical R<sup>2</sup>.
  - Overfit!
- ► Train algorithms on full data, collect more data to evaluate R².

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- ► Train algorithms on full data, collect more data to evaluate R<sup>2</sup>.
  - Expensive!

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- ► Train algorithms on full data, see which has largest empirical R<sup>2</sup>.
  - Overfit!
- ► Train algorithms on full data, collect more data to evaluate R<sup>2</sup>.
  - Expensive!
- ► Cross validation!

Consider randomly splitting the data into K different pieces.

$S_1$
$S_2$
$S_3$
$S_4$
$S_5$

Define first training,  $T_1$ , and validation,  $V_1$ , sample.

$V_1$
$T_1$
$\mathrm{T}_1$
$T_1$
$T_1$

Train M algorithms using  $T_1$ .

•  $\psi_{j,m}(T_1)$ , for  $m=1,\ldots,M$ 

Withhold validation sample  $V_1$  from training process.

► As though we did another experiment of size  $|V_1|$ !

Use validation sample to estimate MSE of each algorithm

$$\hat{\mathcal{E}}_{j,m,1}(\psi_{j,m}) = \frac{1}{|V_1|} \sum_{i \in V_1} \{Y_{j,i} - \psi_{j,m}(T_1)(X_i)\}^2.$$

Define second training, T<sub>2</sub>, and validation, V<sub>2</sub>, sample.

$T_2$
$V_2$
$\mathrm{T}_2$
$T_2$
$T_2$

Train M algorithms using  $T_2$ .

 $\blacktriangleright \ \psi_{j,m}(T_2), \text{ for } m=1,\ldots,M$ 

Withhold validation sample  $V_2$  from training process.

► As though we did another experiment of size  $|V_2|$ !

Use validation sample to estimate MSE of each algorithm

$$\hat{\mathcal{E}}_{j,m,2}(\psi_{j,m}) = \frac{1}{|V_2|} \sum_{i \in V_2} \{Y_{j,i} - \psi_{j,m}(T_2)(X_i)\}^2.$$

Continue until each split has been validation once.

$T_3$
$T_3$
$V_3$
$T_3$
$T_3$

Continue until each split has been validation once.

$T_4$
$T_4$
$\mathrm{T}_4$
$V_4$
$T_4$

Continue until each split has been validation once.

$T_5$
$T_5$
$T_5$
$T_5$
$V_5$

#### Cross-validation selector

The cross-validated MSE of algorithm m is

$$\hat{\mathcal{E}}_{j,m}(\psi_{j,m}) = \frac{1}{K} \sum_{k=1}^{K} \hat{\mathcal{E}}_{j,m,k}(\psi_{j,m}) .$$

We call the algorithm  $m^*$  with the lowest MSE the cross-validation selector.

We might use  $\psi_{j,m^*}(F_n)$  as estimate of  $E_0(Y_j \mid X)$ , where  $F_n = \{1, \dots, n\}$ .

#### **Ensemble estimator**

What if  $\psi_{j,1}$  and  $\psi_{j,2}$  capture different features?

Using  $\psi_{j,\rm SL}=0.5\psi_{j,1}+0.5\psi_{j,2}$  might be better than  $\psi_{j,1}$  and  $\psi_{j,2}$  alone.

More generally, consider an ensemble prediction function

$$\psi_{\rm j,SL} = \sum_{m=1}^{\rm M} \alpha_{\rm j,m} \psi_{\rm j,m} \;,\; \alpha_{\rm j,m} \geq 0 \; {\rm for \; all} \; m \;,\; {\rm and} \sum_{m=1}^{\rm M} \alpha_{\rm j,m} = 1 \;.$$

Easy to find  $\alpha_j$  that minimizes cross-validated MSE.

# Super learner

Stacked regression originally proposed in Breiman, 1996.

Referred to as a super learner due to oracle inequality results (van der Laan and Dudoit, 2003).

- ► MSE of the super learner is asymptotically equivalent to the oracle estimator.
- ► Even when one considers many estimators.
- ► Often seen to have good finite-sample performance (van der Laan et al, 2007).

Proving new oracle results is an open area of research!

► Recent work on Big Data oracle inequalities.

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# Statistical goal

Difficult optimization problem,

$$\psi_{0,\omega_0} = \operatorname{argmax}_{\psi} R_{0,\omega_0}^2(\psi)$$
  
$$\omega_0 = \operatorname{argmax}_{\omega} R_{0,\omega}^2(\psi_{0,\omega}) ,$$

made easier because

$$\psi_{0,\omega} = \mathcal{E}_0\left(\sum_{j=1}^J \omega_j \mathcal{Y}_j \mid \mathcal{X}\right) = \sum_{j=1}^J \omega_j \mathcal{E}_0(\mathcal{Y}_j \mid \mathcal{X}) .$$

For any  $\omega$ , combine super learners to estimate  $\psi_{0,\omega}$ ,

$$\psi_{n,\omega} = \sum_{j=1}^{J} \omega_j \psi_{j,SL}(F_n) .$$

How do we estimate  $\omega_0$ ?

#### How do we estimate $\omega_0$ ?

► Maximize empirical R<sup>2</sup> over weights.

$$\omega_n = \operatorname{argmax}_{\omega} \left( 1 - \frac{\frac{1}{n} \sum_{i=1}^{n} \{ Y_{\omega,i} - \psi_{n,\omega}(X_i) \}^2}{\frac{1}{n} \sum_{i=1}^{n} \{ Y_{\omega,i} - \bar{Y}_{\omega} \}^2} \right)$$

How do we estimate  $\omega_0$ ?

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overfit!

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$$\omega_n = \operatorname{argmax}_{\omega} \left( 1 - \frac{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \psi_{n,\omega}(X_i)\}^2}{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \bar{Y}_{\omega}\}^2} \right)$$

- overfit!
- ► Collect more data, maximize R<sup>2</sup> on new data.

# Estimating optimal weights

#### How do we estimate $\omega_0$ ?

► Maximize empirical R<sup>2</sup> over weights.

$$\omega_n = \operatorname{argmax}_{\omega} \left( 1 - \frac{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \psi_{n,\omega}(X_i)\}^2}{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \bar{Y}_{\omega}\}^2} \right)$$

- overfit!
- ► Collect more data, maximize R<sup>2</sup> on new data.
  - expensive!

# Estimating optimal weights

How do we estimate  $\omega_0$ ?

► Maximize empirical R<sup>2</sup> over weights.

$$\omega_n = \operatorname{argmax}_{\omega} \left( 1 - \frac{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \psi_{n,\omega}(X_i)\}^2}{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \bar{Y}_{\omega}\}^2} \right)$$

- overfit!
- ► Collect more data, maximize R<sup>2</sup> on new data.
  - expensive!
- ► Cross validation!

Define first training,  $T_1$ , and validation,  $V_1$ , sample.

7	$V_1$
r	$\Gamma_1$

## **Cross validation**

Train J super learners using  $T_1$ .

 $\blacktriangleright \ \psi_{j,SL}(T_1), \text{ for } j=1,\ldots,J$ 

For any  $\omega$ , we can construct combined super learner

$$\psi_{\omega,SL}(T_1) = \sum_{j=1}^{J} \omega_j \psi_{j,SL}(T_1) .$$

Withhold validation sample  $V_1$  from super learner fitting.

▶ As though we did another experiment of size  $|V_1|$ !

For any  $\omega$ , use validation sample to estimate MSE

$$\hat{\mathcal{E}}_{\omega,1}(\psi_{\omega,SL}) = \frac{1}{|V_1|} \sum_{i \in V_1} \{Y_{\omega,i} - \psi_{\omega,SL}(T_1)(X_i)\}^2.$$

Define second training,  $T_2$ , and validation,  $V_2$ , sample.

$T_2$
$V_2$
$T_2$
$T_2$
$T_2$

Train J super learners using  $T_2$ .

 $\blacktriangleright \ \psi_{j,\mathrm{SL}}(\mathrm{T}_2), \, \mathrm{for} \, j=1,\ldots,\mathrm{J}$ 

For any  $\omega$ , we can construct combined super learner

$$\psi_{\omega,SL}(T_2) = \sum_{j=1}^{J} \omega_j \psi_{j,SL}(T_2) .$$

Withhold validation sample  $V_2$  from super learner fitting.

lacktriangle As though we did another experiment of size  $|V_2|!$ 

For any  $\omega$ , use validation sample to estimate MSE

$$\hat{\mathcal{E}}_{\omega,2}(\psi_{\omega,SL}) = \frac{1}{|V_2|} \sum_{i \in V_2} \{Y_{\omega,i} - \psi_{\omega,SL}(T_2)(X_i)\}^2.$$

Continue until each split has been used once.

$T_3$
$T_3$
$V_3$
$T_3$
$T_3$

Continue until each split has been used once.

$T_4$
$\mathrm{T}_4$
$\mathrm{T}_4$
$V_4$
$T_4$

Continue until each split has been used once.

$T_5$	
$T_5$	
$T_5$	
$T_5$	
$V_5$	
$T_5$	

## **Estimating weights**

For any  $\omega$ , we have cross-validated estimate of  $\mathbb{R}^2$ ,

$$R_{n,\omega}^{2}(\psi_{\omega,SL}) = 1 - \frac{\frac{1}{K} \sum_{k=1}^{K} \hat{\mathcal{E}}_{\omega,k}(\psi_{\omega,SL})}{\frac{1}{n} \sum_{i=1}^{n} \{Y_{\omega,i} - \bar{Y}_{\omega}\}^{2}}.$$

Estimate of optimal weights is

$$\omega_n = \operatorname{argmax}_{\omega} R_{n,\omega}^2(\psi_{\omega,SL}) .$$

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Conclusions and future directions

How accurate is  $\psi_{n,\omega_n}$  in predicting  $Y_{\omega_n}$ ?

► Report  $R_{n,\omega_n}^2(\psi_{\omega_n,SL})$ , call it a day.

- ► Report  $R_{n,\omega_n}^2(\psi_{\omega_n,SL})$ , call it a day.
  - overfit!

- ► Report  $R_{n,\omega_n}^2(\psi_{\omega_n,SL})$ , call it a day.
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- ► Collect more data to evaluate predictions.

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  - expensive!

- ► Report  $R_{n,\omega_n}^2(\psi_{\omega_n,SL})$ , call it a day.
  - overfit!
- ► Collect more data to evaluate predictions.
  - expensive!
- ► More cross validation!!!

# Doubly nested cross validation

Pictures omitted for the sanity of audience.

Cross-validate the entire procedure to estimate performance.

- ► Compute  $\omega_n$  and  $\psi_{n,\omega_n}$  in training sample
- ► Estimate R<sup>2</sup> in validation sample
- ► Average over splits

Formally, we define  $\omega : \mathcal{S} \to \Omega$  and  $\psi_{\omega, SL} : \mathcal{S} \to \Psi$ .

The cross-validated R<sup>2</sup> estimate is

$$R_n^2(\omega_n,\psi_{n,\omega_n}) = 1 - \frac{\sum_{k=1}^K \frac{1}{|V_k|} \sum_{i \in V_k} \{Y_{\omega(T_k),i} - \psi_{\omega(T_k),SL}(X_i)\}^2}{\sum_{k=1}^K \frac{1}{|V_k|} \sum_{i \in V_k} \{Y_{\omega(T_k),i} - \bar{Y}_{\omega(T_k)}\}^2} \ .$$

## Inference for predictive performance

In spite of the highly adaptive estimation procedure,

$$n^{1/2}\left\{R_n^2(\omega_n,\psi_{n,\omega_n})-R_0^2(\omega_n,\psi_{n,\omega_n})\right\} \to \text{Normal}(0,\sigma^2)$$
.

A sufficient condition is that  $\omega_0$  is unique.

▶ Possible to relax this condition (Luedtke et al, 2016).

Variance derived via delta method for influence functions.

► Consistently estimated with closed-form estimator.

Variance estimates used to construct closed-form confidence intervals and hypotheses tests.

► Machine learning with inference!

## Variable importance

Define  $R_0^2(\omega_n^d, \psi_{n,\omega_n}^d)$  as the true  $R^2$  of the estimated super learner that leaves out  $X_d, d = 1, ..., D$ .

The "importance" of X<sub>d</sub> could be quantified by

$$\Delta_{0n}^d = R_0^2(\omega_n, \psi_{n,\omega_n}) - R_0^2(\omega_n^d, \psi_{n,\omega_n}^d) .$$

How much did  $X_d$  improve predictions of combined outcome?

► Similar to random forest variable importance measures

# Variable importance

Variable importance can be estimated as

$$\Delta_n^d = R_n^2(\omega_n, \psi_{n,\omega_n}) - R_n^2(\omega_n^d, \psi_{n,\omega_n}^d) .$$

We can still establish

$$n^{1/2}(\Delta_n^d - \Delta_{0n}^d) \to \text{Normal}(0, \sigma_d^2)$$
.

Variance can again be derived using delta method for influence functions.

Did  $X_d$  significantly improve predictions of combined outcome?

## **Outline**

#### Motivation

### Defining target parameters

- Measuring accuracy of predictions
- ► Optimal predictor and weights

#### Estimation

- Super learning
- ► Estimating weights

### **Evaluating predictions**

- ► Estimation and inference
- ► Variable importance

#### Simulation

### Data Analysis

Conclusions and future directions

### Simulation

#### Covariates:

$$X_1, \dots, X_6 \sim Uniform(0, 4), X_7, \dots, X_9 \sim Bernoulli(0.5)$$

#### **Outcomes:**

$$\begin{split} Y_1 &= X_1 + 2X_2 + 4X_3 + X_7 + 2X_8 + 4X_9 + 2X_4 + \epsilon_1 \;, \\ Y_2 &= X_1 + 2X_2 + 4X_3 + X_7 + 2X_8 + 4X_9 + 2X_5 + \epsilon_2 \;, \text{ and} \\ Y_3 &= X_1 + 2X_2 + 4X_3 + X_7 + 2X_8 + 4X_9 + 2X_6 + \epsilon_3 \;, \end{split}$$

with  $\epsilon_j \sim \text{Normal}(0, 5^2), j = 1, 2, 3$ .

### True parameters:

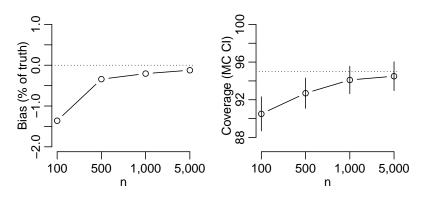
$$\omega_0=(\overline{\frac{1}{3}},\overline{\frac{1}{3}},\overline{\frac{1}{3}})$$
,  $R_{0,\omega_0}^2=0.80$ ,  $\Delta_0^2=0.12$ .

### Super learner:

intercept only, main terms, and stepwise linear model.

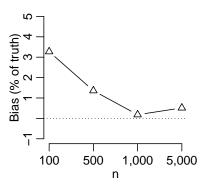
## Simulation

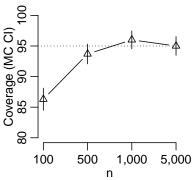
Bias and coverage for  $R_{n,\omega_n}^2$  (1000 replications).



## Simulation

Bias and coverage for  $\Delta_n^2$  (1000 replications).





### **Outline**

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Conclusions and future directions

## Data analysis

Can we use neonatal information to predict neurocognitive outcomes later in life?

► Early identification of at-risk children.

What covariates are important for making predictions?

► Informs what information to collect to screen children.

## **Motivation: PROBIT**

The Promotion of Breastfeeding Intervention Trial enrolled pregnant mothers in 1996–97.

Variables
breastfeeding encouraged
household animals
age, height, weight, education, siblings
employment status
gestational age, Apgar score
WAZ, HAZ, HCAZ (0,1,2,3,6,9,12 months)
mother smoked during pregnancy, mother
drank during pregnancy
Matrix, Block, Vocabulary, Similarities
-

### **Future directions**

Effect estimation for discovery in high dimensions.

- ► Maximize effect instead of R<sup>2</sup>?
- ► Cross-validated TMLE (Zheng and van der Laan, 2010)

The method could be extended to binary outcomes, other performance metrics, and dependent data.

Nonlinear combinations of outcomes are also of interest.

 Alternating conditional expectations (Breiman and Friedman, 1985)

### Software

### R packages:

### r2weight

Available on GitHub: https://github.com/benkeser/r2weight

SuperLearner (Polley et al, 2016)

► Demonstration - http://benkeser.github.io/sllecture/

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# **Oracle Inequality**

Let  $L: \Psi \times \mathcal{O} \to \mathbb{R}$  be a loss function for  $\psi_0$  in the sense that

$$\psi_0 = \text{argmin}_{\psi \in \Psi} E_0\{L(\psi)(O)\} \ .$$

Define  $d_0(\psi, \psi_0) := \mathbb{E}_0\{L(\psi)(O) - L(\psi_0)(O)\}$  and let p be the proportion of observations in the validation sample. Assume

- 1.  $\psi_{n,m} \in \Psi$  with probability tending to one for m = 1, ..., M.
- 2. For some  $C_0 < \infty$ ,  $\sup_{\psi \in \Psi} d_0(\psi, \psi_0) < C_0$  almost surely.
- 3. For some  $C_1 < \infty$ ,  $E_0\{L(\psi)(O) L(\psi_0)(O) d_0(\psi, \psi_0)\}^2 \le C_1 d_0(\psi, \psi_0)$  for all  $\psi \in \Psi$ .

For every  $\lambda > 0$  and  $C(\lambda) := \frac{2}{3}(1+\lambda)^2(C_0 + C_1)$ ,

$$\begin{split} & \mathbb{E}_{\mathbb{B}_{n}} \big\{ d_{0}(\psi_{n, \text{SL}}, \psi_{0}) \big\} \\ & \leq (1 + 2\lambda) \mathbb{E}_{\mathbb{B}_{n}} \big\{ d_{0}(\psi_{n, \text{OR}}, \psi_{0}) \big\} + 2\mathbb{C}(\lambda) \bigg( \frac{1 + \log K(n)}{np} \bigg) \end{split}$$

# Asymptotics

The cross-validated  $R^2$  estimate is

$$R_n^2(\omega_n, \psi_{n,\omega_n}) = 1 - \frac{\sum_{k=1}^K \frac{1}{|V_k|} \sum_{i \in V_k} \{Y_{\omega(T_k),i} - \psi_{\omega(T_k),SL}(X_i)\}^2}{\sum_{k=1}^K \frac{1}{|V_k|} \sum_{i \in V_k} \{Y_{\omega(T_k),i} - \bar{Y}_{\omega(T_k)}\}^2}$$
$$= 1 - \frac{\theta_{1,n}}{\theta_{2,n}}$$

For  $k = 1, \dots, K$ , define

$$\begin{split} &D_{0n,k}(\psi_{\omega})(O) \\ &:= \{Y_{\omega(T_k)} - \psi_{\omega(T_k)}(T_k)(X)\}^2 - E_0\left[\{Y_{\omega(T_k),i} - \psi_{\omega(T_k),SL}(X_i)\}^2\right] \;, \\ &D_{0n,k}(\bar{Y}_{\omega})(O) \\ &:= \{Y_{\omega(T_k)} - \bar{Y}_{\omega(T_k)}\}^2 - E_0\left[\{Y_{\omega(T_k),i} - \bar{Y}_{\omega(T_k)}(X_i)\}^2\right] \;. \end{split}$$

# Asymptotics

$$n^{1/2}(\theta_{1,n} - \theta_{1,0}) \to \text{Normal}(0, \sigma_1^2), \text{ with }$$

$$\sigma_1^2 = \frac{1}{K} \sum_{k=1}^K E_0 \{ D_{0n,k}(\psi_{\omega_0})(O)^2 \} .$$

 $n^{1/2}(\theta_{2,n} - \theta_{2,0}) \to \text{Normal}(0, \sigma_2^2), \text{ with }$ 

$$\sigma_2^2 = \frac{1}{K} \sum_{k=1}^K E_0 \{ D_{0n,k}(\bar{Y}_{\omega_0})(O)^2 \} .$$

Let  $D_{0n,k}(\psi_{\omega}, \bar{Y}_{\omega}) = (D_{0n,k}(\psi_{\omega}), D_{0n,k}(\bar{Y}_{\omega}))^{T}$ ,  $g(\theta) = \log(\theta_{1}/\theta_{2})$ , and  $\nabla g(\theta) = (1/\theta_{1}, -1/\theta_{2})^{T}$ .

# Asymptotics

We have  $n^{1/2}\{g(\theta_n) - g(\theta_0)\} \to \text{Normal}(0, \sigma_3^2)$ , where  $\sigma_3^2$  is

$$\nabla g(\theta_0)^{\mathrm{T}} \frac{1}{K} \sum_{k=1}^{K} \mathrm{E}_0 \{ \mathrm{D}_{0n,k}(\psi_{\omega_0}, \bar{\mathrm{Y}}_{\omega_0})(\mathrm{O}) \mathrm{D}_{0n,k}(\psi_{\omega_0}, \bar{\mathrm{Y}}_{\omega_0})(\mathrm{O})^{\mathrm{T}} \} \nabla g(\theta_0) .$$

## Canonical correlation

Let 
$$Y_{\alpha} = \sum_{j=1}^{J} \alpha_j Y_j$$
 and  $X_{\beta} = \sum_{d=1}^{D} \beta_d X_d$ .

The first-order canonical variate of X and Y is found by maximizing

$$\frac{E_0\{(Y_{\alpha}-\mu_{0,\alpha})(X_{\beta}-\mu_{0,\beta})\}}{E_0\{(Y_{\alpha}-\mu_{0,\alpha})^2\}E_0\{(X_{\beta}-\mu_{0,\beta})^2\}}$$

over  $\alpha$  and  $\beta$  under constraint that variances equal one.

The canonical correlation is the correlation between  $X_{\beta,0}$  and  $Y_{\alpha_0}$ .

### Canonical correlation

If  $\psi_{0,j} = X_{\beta}$  for all j = 1, ..., J, the optimal  $\mathbb{R}^2$  equals the squared first-order canonical correlation.

To illustrate difference, consider

- ►  $X_d$  ~ Normal(0,1), d = 1, 2
- ►  $Y_j = X_j^2 \text{ for } j = 1, 2$

Canonical correlation measures linear association between X and Y.

► Canonical correlation equals zero.

Optimal  $\mathbb{R}^2$  measures how well we predict Y using X.

► Optimal R<sup>2</sup> equals one.